

AN ANALYSIS OF THE EFFECT OF INCOME ROUNDING  
IN THE CURRENT POPULATION SURVEY

by

Joseph J. Knott

Social and Economic Statistics Administration, Bureau of the Census

This paper analyzes a form of income reporting bias caused by income sample survey respondents reporting income in rounded or approximate amounts.

This paper makes 7 major points:

1. An income rounding bias exists.
2. Responses from "proxy" respondents show more bias than responses from persons who report their own income to the interviewer.
3. The propensity to round is directly related to income.
4. Female persons have more tendency to round income data than male persons.
5. The bias is found at all levels of income aggregation with respect to persons income.
6. The bias may affect the usefulness of the CPS income data for some very specialized uses.
7. There are several changes in procedures which will reduce the bias.

#### Income Rounding Bias

It is well known that in any sample survey involving voluntary responses, such as the CPS, respondents do not always report exact income data. In fact, interviewers are specifically instructed to accept a reasonable estimate rather than report the income item as a non-response or "NA" (not available). This estimation bias has never been to my knowledge specifically quantified, although its presence has been recognized by many researchers. As a result of this bias, when income is tabulated by intervals which are less than \$1,000 wide, the interval containing an exact thousand dollar amount or an exact \$500 amount shows an extraordinarily high reporting frequency. This bias is most easily illustrated by a bar graph by \$100 intervals. (see graph 1) this rather peculiar distribution is at variance with any of the usual income size distribution models (normal, lognormal, log-log).

#### Proxy vs Self Reporting

In the CPS interview process interviewers are not required to interview every member of the household, but may accept information from a responsible and knowledgeable "proxy" respondent. There were 40,046 proxy respondents or 43.0 percent of the 93,193 persons 14 years old and

over who responded to the March, 1970 CPS survey. As the horizontal percentages in table 1 show, over half of the responses for male persons were provided by proxy respondents, but only about 30 percent of the responses for female persons were made by proxy respondents.

The percent of total persons reporting income in the \$n,000 - \$n,099 (n=1 to .14) intervals who had proxy respondents is higher in all 14 income intervals than the percent reporting for themselves who reported in that interval. Above \$4,000, the percent is higher than 25 percent in all intervals tabulated (see table 2).

#### Propensity to Round Income Amounts

The propensity to round income amounts is in general directly related to the amount of income in the \$1,000 to \$15,000 range. As table 2 shows, the propensity to report income in the lowest \$100 interval containing the exact thousand dollar amount increases as income increases. The tendency to round is very pronounced in the \$10,000, \$11,000, and \$12,000 intervals. In several of the \$1,000 intervals over 50 percent of the units reported the \$n,000 to \$n,099 interval (see table 2).

#### Reporting by Male vs Female Persons

In general, female self respondents have more of a tendency to report in the \$n,000 - \$n,099 than male self respondents. In only 4 out of the 14 intervals are the percent of male higher or equal to the percent reporting in the lowest \$100 of the \$1,000 interval. (see table 2).

#### Levels of Aggregation

The income rounding bias is present at all levels of aggregation, from a single type of income for a person to total family income. As Table 3 shows, the percent of persons reporting in the lowest \$100 of an income class is higher for the individual income types than for total money income. This is true for both male and female respondents. The presence of the bias in total family income has been reported by this author in another paper. <sup>1/</sup>

The degree in which the bias is present in persons total money income for persons and for total family income is the surprising aspect of the bias. It had been the assumption of this author and other researchers that as the eight individual income types were aggregated to total persons income and as persons income was aggregated to total family income, that the effect of the income bias would be greatly reduced if not eliminated. However, it seems clear that the degree of the income bias exceeds

the expected amount.

### Impact of the Rounding Bias

In general the rounding bias does not affect the March income data greatly. The March income supplement to the CPS is designed to collect general statistics on the National level. In this sense, I doubt if the cost of collecting more precise data would be justified by the benefit derived. However, for some specialized researchers, this may not be the case. More and more the March CPS file is being used as a microdata file and for these users the rounding bias may cause problems. The possible impact of the income rounding bias on the analysis of the low income population and on econometric modeling is pointed out below. Also the impact on income interval means and interpolation is discussed.

### Economic Models

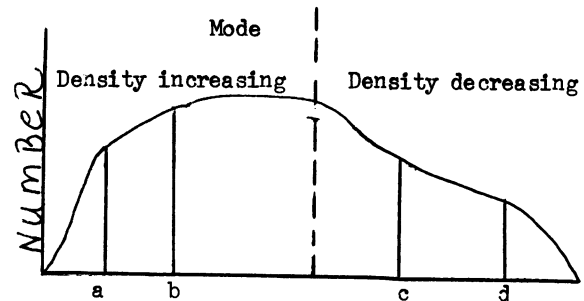
One use of the CPS March file is as input to economic models to evaluate the cost of effects of various programs. This is usually done by assigning different eligibility is dependent on income, the income reporting bias could be reflected by unexpected jumps in the number of eligibles as the qualification income threshold level is increased to include these discontinuous "lumps" at thousand and five hundred dollar levels. I doubt that the income reporting bias would seriously effect the overall conclusion, but for specialized types of analysis it may.

### Possible Impact on the Analysing Low Income Population

Since the low income thresholds are adjusted by the Consumer Price Index (CPI) each year, the rounding bias could cause aberrations in the numbers of families classified as low income. For example, an increase in a low income threshold from 3,899 to 3,999 will increase the number of low income families less than an increase from 3,999 to 4,099. However, this aspect of the income rounding bias requires more analysis before its effect on changing the number of lower income families can be understood.

### Income Interval Means

Another consequence of rounding reported income values is that the income class interval means are depressed. It is the "usual" practice to tabulate income in \$1,000 or \$500 intervals except for high income values. For the intervals close to the class containing the mode income value, one can assume that the actual income values are distributed in the manner shown below. For the income class intervals below the mode (class ab for example), it is expected that the class means would fall above the midpoint ( $\bar{ab} < \frac{b+a}{2}$ ). For the income class intervals above the mode (cd for example), it is expected that the class mean would fall below the midpoint ( $\bar{cd} > \frac{c+d}{2}$ ).



However, this is not the case for class interval below the mode. As table 4 clearly shows the class intervals below the median (assumed to be close to the mode) the tabulated class falls well below the midpoint in the March Current Population Survey (CPS) income data. The Internal Revenue Service class means (table 4) falls slightly below the class midpoint but are much closer to the expected value than the CPS. The IRS data may also indicate some slight tendency for money to be paid in round amounts.

### Interpolation for Median Income

It is standard practice to assume a linear distribution within a income class for the computations of the median income value. As can be seen from the \$100 interval bar graph this appears to be wrong assumption (see graph 1-3). A negative exponential function ( $y = \ln e^{-x}$ ) would be a better fit for the observed \$100 interval values within a \$1000 income class. However, since I believe the observed reporting pattern to be the result of a rounding bias, the assumption of a linear distribution may more closely conform with the true income distribution in the real world.

### Reduction of the Income Bias

The income rounding bias could be reduced by several modifications in the instructions covering the field work on collecting the data. These are:

1. To the extent possible, interview the respondent personally.
2. To the extent possible, encourage the respondent to consult records for exact amounts rather than relying on memory.
3. Interviewers should be less willing to accept estimates.

While these modifications will reduce the income rounding bias they will also increase cost of field work and probably increase the income non-response rate. These tradeoffs would have to be considered implementing the proposed modifications.

### Tabulation Modification

The importance of the bias can be reduced by tabulating data in \$1000 intervals which start

with the \$500 level. For example, under \$500, \$500 to \$1499, \$1500 to \$2499, \$2500 to \$3499, etc. Also it must be recognized that tabulation by intervals less than \$1000 in size will result in some distortion of the data.

#### Conclusions

The data presented in this paper clearly show the presence of an income rounding bias. The hypothesis that the bias is caused by respondents reporting in exact thousand dollar amount and to a lesser extent in exact five hundred dollar amount has been established but not proven. Further investigation of this subject is being planned. This will require tabulation of persons reporting exactly an exact thousand dollar amount and five hundred dollar amount.

1/ "The Index of Income Concentration in the 1970 Census of Population and Housing", Joseph J. Knott, ASA Proceedings of the Social Science Section 1971.

Table 1.--Respondents 14 Years Old and Over by Sex and Self vs. Proxy Reporting Status in 1969

	Total	Self	Proxy
Total	93,192	53,146	40,046
Male	42,619	17,614	25,005
Female	50,573	35,532	15,041
Horizontal Percent			
Total	100.0	57.0	43.0
Male	100.0	41.3	58.7
Female	100.0	70.3	29.7
Vertical Percent			
Total	100.0	100.0	100.0
Male	45.7	33.1	62.4
Female	54.3	66.9	37.6

Table 2.--Number of Units Reporting Total Money in the \$n,000 - \$n,099 Interval as a Percent of the Units in the \$n,000 - \$n,099 Interval by Self and Proxy

Percent of Units in \$n,000 - \$n,099 Intervals						
Total Money Income	Total		Male		Female	
	Self	Proxy	Self	Proxy	Self	Proxy
\$1,000 - \$1,999	14.0	16.1	11.6	15.1	15.1	18.4
\$2,000 - \$2,999	16.8	20.1	14.3	18.7	18.1	22.2
\$3,000 - \$3,999	18.7	20.7	17.4	19.3	19.4	22.7
\$4,000 - \$4,999	20.2	22.0	18.9	20.7	20.9	24.1
\$5,000 - \$5,999	22.1	25.3	21.1	22.9	23.0	31.3
\$6,000 - \$6,999	23.6	26.1	21.7	24.7	25.7	32.8
\$7,000 - \$7,999	18.9	24.2	19.0	22.8	18.7	35.1
\$8,000 - \$8,999	24.0	27.7	23.4	27.1	25.6	35.2
\$9,000 - \$9,999	22.7	28.5	22.5	27.9	23.8	40.7
\$10,000 - \$10,999	29.8	35.3	29.8	35.4	29.8	54.2
\$11,000 - \$11,999	27.5	31.1	28.3	30.7	23.2	42.4
\$12,000 - \$12,999	32.7	40.3	32.1	39.9	36.2	52.0
\$13,000 - \$13,999	26.6	31.6	26.0	30.9	31.2	58.3
\$14,000 - \$14,999	26.7	28.5	27.0	28.4	24.3	40.0

NOTE: See Table 3 for details.

TABLE 3.-- Percent of Persons 14 Years Old or Over who Reported Income of The Specified Type and Reported in the Lowest \$100 of The Specified \$1000 Interval in 1969 by SEX and Reporting STATUS (Self vs. Proxy)

INCOME LEVEL \$1,000	PERCENT OF UNITS IN THE \$1,000 - \$1,099 INTERVAL																	
	SELF									PROXY								
	TOTAL MONEY INCOME	WAGES AND SALARIES	NONFARM SELF- EMPLOYED	FARM SELF- EMPLOYED	SOCIAL SECURITY + RAILROAD RETIREMENT	DIVIDENDS INTEREST NET RENT ECT	PUBLIC ASSISTANCE	PUBLIC TRANSFERS	PRIVATE PENSIONS ANNUITIES ALIENAGE, ECT	TOTAL MONEY INCOME	WAGES AND SALARIES	NONFARM SELF- EMPLOYED	FARM SELF- EMPLOYED	SOCIAL SECURITY + RAILROAD RETIREMENT	DIVIDENDS INTEREST NET RENT ECT	PUBLIC ASSISTANCE	PUBLIC TRANSFERS	PRIVATE PENSIONS ANNUITIES
MALE																		
1000 - 1999	11.6	12.8	25.6	22.0	8.1	24.3	25.6	24.6	14.6	15.1	17.1	37.1	30.4	9.8	24.3	29.9	16.9	18.1
2000 - 2999	14.3	22.1	38.3	37.4	12.4	29.1	13.3	10.2	13.3	18.7	22.9	38.1	35.8	14.6	24.9	11.1	12.7	18.2
3000 - 3999	17.4	21.9	40.0	37.5	20.3	42.3	17.6	15.1	26.9	19.3	20.8	54.4	43.1	12.5	39.0	9.1	28.3	40.6
4000 - 4999	18.9	23.4	42.1	50.0	20.0	37.2	16.7	20.3	25.0	20.7	23.3	38.0	48.9	7.7	37.7	20.1	15.4	42.1
5000 - 5999	21.1	24.1	61.1	44.6	(B)	52.3	-	15.0	45.2	22.9	26.5	48.1	40.0	(B)	50.0	(B)	20.0	44.4
6000 - 6999	21.1	23.9	62.0	40.5	(B)	54.5	(B)	36.4	36.4	24.7	27.6	56.6	36.7	(B)	40.0	(B)	33.3	(B)
7000 - 7999	19.0	21.4	43.8	65.2	(B)	28.6	-	22.2	33.3	22.8	25.2	49.3	51.9	-	25.0	(B)	28.6	(B)
8000 - 8999	23.4	27.3	55.8	45.8	-	40.0	(B)	37.5	40.0	27.1	32.7	69.4	36.8	-	37.5	-	14.3	(B)
9000 - 9999	22.5	28.7	62.2	50.0	(B)	46.7	-	36.4	(B)	27.9	32.0	61.5	66.7	-	14.3	-	57.1	(B)
10000 - 10999	29.8	39.2	75.7	72.7	-	68.4	-	(B)	40.0	35.4	43.3	83.8	76.5	-	56.3	-	(B)	(B)
11000 - 11999	28.3	42.3	72.7	83.3	-	57.1	-	(B)	(B)	30.7	41.3	54.5	50.0	-	(B)	-	-	-
12000 - 12999	32.1	52.4	84.6	77.8	-	50.0	-	40.0	(B)	39.9	54.9	76.8	83.3	-	(B)	-	(B)	-
13000 - 13999	26.0	45.5	72.2	60.0	-	33.0	-	-	(B)	30.9	42.5	46.2	(B)	-	(B)	-	(B)	(B)
14000 - 14999	27.0	39.0	70.0	33.3	-	(B)	-	-	-	28.4	38.6	75.0	(B)	-	0.0	-	-	(B)
FEMALE																		
1000 - 1999	15.1	16.9	24.4	25.7	16.5	22.3	22.2	17.0	14.1	18.4	18.2	31.8	(B)	21.8	24.6	26.4	21.1	20.2
2000 - 2999	18.1	19.9	37.7	30.0	13.4	27.1	14.8	20.5	17.5	22.2	22.5	29.4	(B)	36.4	32.4	21.4	11.1	22.2
3000 - 3999	19.4	21.0	32.6	28.6	13.2	30.3	15.8	14.3	24.5	22.7	23.7	46.2	(B)	(B)	25.0	0.0	(B)	66.7
4000 - 4999	20.9	23.6	41.4	(B)	26.5	28.1	8.0	29.2	19.5	24.1	25.6	12.5	-	(B)	33.3	(B)	(B)	33.3
5000 - 5999	23.0	26.1	47.2	42.9	(B)	37.9	(B)	30.0	29.4	31.3	32.9	(B)	-	-	66.7	(B)	(B)	-
6000 - 6999	25.7	27.9	45.5	33.3	(B)	37.5	(B)	(B)	50.0	32.8	34.2	(B)	(B)	-	(B)	-	(B)	(B)
7000 - 7999	18.7	19.8	38.5	-	(B)	30.8	(B)	(B)	0.0	35.1	37.7	-	-	-	62.5	-	-	(B)
8000 - 8999	25.6	31.0	40.0	-	-	69.2	-	(B)	(B)	35.2	37.0	(B)	-	(B)	(B)	-	-	-
9000 - 9999	23.8	30.6	(B)	-	(B)	36.4	-	(B)	40.0	40.7	37.9	(B)	-	(B)	(B)	-	(B)	(B)
10000 - 10999	29.8	33.3	63.6	(B)	-	81.8	-	-	(B)	54.2	57.1	(B)	-	-	(B)	-	-	(B)
11000 - 11999	23.2	36.9	(B)	-	-	40.0	-	(B)	(B)	42.4	57.7	(B)	-	-	(B)	-	-	-
12000 - 12999	36.2	49.3	(B)	-	-	80.0	-	-	(B)	52.0	55.6	(B)	-	-	(B)	-	-	-
13000 - 13999	31.2	51.7	(B)	-	-	16.7	-	-	(B)	58.3	60.0	(B)	-	-	-	-	-	-
14000 - 14999	24.3	35.0	(B)	-	-	-	-	-	-	40.0	66.7	-	-	-	-	-	-	-

(B) - Base less than 5 - NONE

Percent =  $\frac{\text{number } \$1,000 - \$1,099}{\text{number } \$1,000 - \$1,999} \times 100$

TABLE A.--MEAN INCOME of SPECIFIED TYPE by SIZE CLASS for MALES 14 years old and over Reporting for Self in 1969 and ALL INDIVIDUAL INCOME TAX RETURNS in 1968

INCOME of SPECIFIED TYPE	MEAN INCOME OF SPECIFIED TYPE														
	CPS MALES REPORTING FOR SELF WITH SPECIFIED TYPE OF INCOME									ALL RETURNS WITH SPECIFIED TYPE OF INCOME					
	TOTAL	WAGES and Salaries	NONFARM SELF- EMPLOYED	FARM SELF- EMPLOYED	SOCIAL SECURITY RAILROAD RETIREMENT	DIVIDENDS INTEREST NET RENT Etc.	PUBLIC ASSISTANCE	PUBLIC TRANSFERS	PRIVATE PENSIONS, ANNUITIES, Etc.	ADJUSTED GROSS INCOME	WAGES and Salaries	BUSINESS or PROFESSIONAL NET PROFIT	FARM NET PROFIT	DIVIDENDS	INTEREST
TOTAL	NA	NA	NA	NA	NA	NA	NA	NA	NA	\$ 8,783	\$6,839	\$6,403	\$2,993	\$1,355	\$540
Under \$1000	NA	NA	NA	NA	NA	NA	NA	NA	NA	938	502	449	413	191	196
\$1000 TO 1999	\$ 1475	\$ 1438	\$ 1393	\$ 1378	\$ 1486	\$ 1332	\$ 1373	\$ 1331	\$ 1436	1483	1472	1462	1434	1402	1408
2000 TO 2999	2465	2403	2324	2313	2395	2330	2470	2476	2416	2492	2486	2473	2468	2446	2433
3000 TO 3999	3452	3424	3296	3325	3395	3283	3364	3451	3416	3502	3498	3478	3464	3450	3433
4000 TO 4999	4433	4426	4287	4273	4402	4315	4216	4457	4401	4497	4493	4474	4484	4483	4437
5000 TO 5999	5410	5386	5180	5248	(B)	5234	-	5392	5230	5500	5493	5471	5483		
6000 TO 6999	6424	6408	6204	6354	(B)	6256	(B)	6341	6367	6500	6413	6476	6504		
7000 TO 7999	7432	7410	7277	7185	-	7391	-	7391	7375	7500	7508	7479	7490	6921	6647
8000 TO 8999	8390	8366	8210	8282	(B)	8313	(B)	8356	8434	8489	8495	8476	8447		
9000 TO 9999	9401	9367	9195	9295	-	9260	-	9366	(B)	1489	1480	9475	9478		
10000 TO 10999	10358	10295	10105	10109	-	10142	-	(B)	10360						
11000 TO 11999	11376	11279	11104	11033	-	11257	-	(B)	(B)						
12000 TO 12999	12315	12217	12071	12113	-	12210	-	12220	(B)						
13000 TO 13999	13393	13279	13122	13166	-	13266	-	-	(B)	16836	14,690	22,427	16,185	31,971	19,003
14000 TO 14999	14386	14301	14090	(B)	-	(B)	-	-	(B)						
15000 and over	NA	NA	NA	NA	NA	NA	NA	NA	NA						

NA- NOT AVAILABLE (B) - BASE LESS THAN 5

